Determining the Change in Retail Prices of Veal in Turkey by GARCH Method Between 2014-2017 [1]

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[1] This study was presented as oral presentation in the 4th International Congress on Veterinary and Animal Science, 12-15 July 2018, Nevşehir, Turkey

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Abstract
In Turkey, the demand for veal in the red meat market is steadily increasing. In the formation of this demand and price; additional factors such as consumer preferences, government interventions in the red meat market, import decisions, also the implemented policies and subsidies are effective besides the main factors. However, the fact that such practices do not affect the veal prices in the market makes it impossible to ensure price stability and a sustainable market. The analysis of the change in retail prices is important for both producers and consumers so that the validity of this view can be tested. The aim of this study was to analyze the retail prices of minced meat and the veal cubes from the first week of January 2014 to the last week of December 2017 via using the volatility estimation modeling method in the context of the effectiveness of the policies and market interventions applied in the red meat sector implemented in Turkey. Non-stationary series were made stationary by differencing. By developing a Generalised Autoregressive Conditional Heteroskedasticity (GARCH 1,1) model using the series subject to a unit root test, it was found to be the optimal model that best explains the fluctuation of the prices. The results of the estimation show that the retail prices of veal cubes and minced meats fluctuated conspicuously in the said period and that the implemented policies and the market interventions were not adequate to eliminate the instability and uncertainty of the prices.

Keywords: GARCH, Price volatility, Time series, Turkey, Veal

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Türkiyede Perakende Dana Eti Fiyatlarındaki Değişimin Garch Yöntemiyle Belirlenmesi, 2014-2017

Öz

Anahtar sözcükler: GARCH, Fiyat oynaklığı, Zaman serisi, Türkiye, Dana eti

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INTRODUCTION

In Turkey, the demand for red meat is met by sheep, goat, water buffalo and cattle producers, with beef constituting 87.67% of the total red meat production [1]. However, retail beef prices have substantially increased in recent years.

While retail veal prices rose by 130% between 2005-2013, retail mutton prices increased by 154% and the feed prices by 115%. The increases in the prices of red meat were higher than the increases in the CPI (Consumer Price Index), which was 118% in the same period [2]. This drastic increase in the retail prices resulted in approximately 30% decrease in the per capita consumption of red meat [3]. The increase in prices continued in the period from 2014 to 2017 as well, and retail prices of veal rose by 19.54% [4]. The increasing prices of red meat have been a subject of debate both in the media and among the policy-makers in Turkey [2].

The increasing prices both mean a higher level of uncertainty regarding the prices in the next years and poses a significant price risk, which adversely affects producer welfare where there are without risk protection mechanisms in place [3]. The increasing demand for red meat and global pressures on climate change have pushed up wholesale and retail prices of meat products in many parts of the world [5]. Abrupt changes in retail prices usually affect the food security for a part of the population, resulting in changes to the dietary habits [7,8].

Price volatility indicates the range in which the prices may vary in the future [9]. Knowledge of the degree of price volatility is very important in the sense that it would allow policy-makers to formulate appropriate strategies to protect against risks. In primary commodity markets such as the red meat market, production decisions are based on expected prices. When such expectations are not met, producers face the risk of decreasing revenues [10].

Price fluctuations affect all actors in the food supply chain. Volatility decreases the inputs of production and investments, causing a risk for producers who can react [11-13]. Unexpected price increases for consumers who spend a large portion of their income on food have a negative impact on sustainable food consumption [14].

Changes in prices of red meat products are a significant economic risk that affects the level of production. Thus, the research on the changes in retail prices of meat in Turkey is important to understand how to equilibrate the market.

This study is intended to analyse the average retail prices of minced meats and veal cubes under the government policies and interventions in the red meat sector of Turkey between 2014 and 2017 using the volatility estimation modelling method. The main purpose of this study is contributing the literature on the effectiveness of the policies and market interventions applied in the red meat sector. In the long run, this study also aims to provide significant information to the decision and policy makers about the econometric method that they can need.

MATERIAL and METHODS

Data Set

In order to explore the changes in meat prices, the weekly retail prices of minced and veal cubes in Turkey from the first week of 2014 to the last week of 2017 were used as the data for the analysis. The short-term asymmetric effect was determined by considering the time interval to measure the response of price changes. The data sets of the minced meats and veal cubes prices used in the analysis were obtained from the weekly bulletin of the Meat and Milk Board [4].

Analysis Method

The analyses conducted using ARIMA model and ARCH/GARCH models are intended to measure the index uncertainty. In the present study, due to the absence of future predictions, volatility of the price change series has been measured over the current asymmetric effect GARCH model instead of hybrid models. In order to measure the index uncertainty, the ARCH model suggested by Engle [19] for the first time and the GARCH model, which had been developed by Bollerslev [20], were used in the study.

The functioning of the ARCH model can be addressed as follows [21]:

Let's assume a regression model with k explanatory variables.

\[ Y_t = \beta_1 + \beta_2 X_{2t} + \ldots + \beta_k X_{kt} + \varepsilon_t \]

In equation (1), under the assumption that conditional information can be obtained at time (t-1), the error term \( \varepsilon_t \) is normally distributed with zero means and variance \( \varepsilon_t \sim N(0, \alpha_0 + \alpha_1 + \varepsilon_{t-1}^2) \), which can be expressed as \( \varepsilon_t \sim N(0, \alpha_0 + \alpha_1 + \varepsilon_{t-1}^2) \). While the error term with zero mean is one of the assumptions of the method of least squares, the addressing of the error term's variance at time t as a function of the square of the error term at t-1 was introduced by the ARCH model. Additionally, the expression of the error term's variance as \( \alpha_0 + \alpha_1 + \varepsilon_{t-1}^2 \) is called the ARCH(1) process.
Since relatively long lags are used and a fixed lag structure is suggested in applying the ARCH model, some constraints are imposed on the parameters in the conditional variance equation. In an attempt to avoid problems with negative variance parameter estimates, Bollerslev [20] extended the ARCH model and developed a new one that allowed for both more past information and a more flexible lag structure. As the said model was developed not as an alternative to the ARCH model, but as a model intended to improve it, it is called generalised ARCH, or shortly GARCH [22].

Using the conditional variance equation, standard GARCH(p, q) process can be expressed as follows:

\[ \sigma^2_t = \omega + \alpha \varepsilon^2_{t-1} + \beta \sigma^2_{t-1} \]  \[2\]

The equation (2) indicating the conditional variance is a function of the mean (\( \omega \)), ARCH term (\( \varepsilon^2_{t-1} \)) and GARCH term (\( \sigma^2_{t-1} \)). Therefore, \( p \) and \( q \) in GARCH(p, q) denote the ARCH term and GARCH term, respectively. In the ARCH(q) process, conditional variance is a function of only the past sample variances, whereas the GARCH(p, q) process allows lagged conditional variances to enter as well.

When working with time series, one should always pay attention to the stationarity of the data. Time series data of economic and financial variables usually involve trend and seasonality, which may violate the principle of stationary time series [23]. A stationary time series can be defined as mean and variance are independent over time. Where time series are not stationary, the estimated econometric models may produce misleading results. Thus, in the econometric analyses conducted using time series, unit root test (stationarity test) is performed on the time series. Therefore, Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) tests were conducted to test whether the time series had a unit root (stationarity). Results of the unit root test are given in Fig. 1 and Fig. 2 as well as in Table 1 and Table 2.

**Fig. 1 and Fig. 2** present the change in the prices of veal cubes and minced meats over time. It is clear that the variables do not follow a stationary course during the periods under consideration. The veal cubes and minced meats series were rendered stationary after the ADF and PP tests.

**RESULTS**

Decreases and increases in the retail prices of minced meats and veal cubes indicate a significant level of price fluctuation. Following the increases in October 2014 and May 2015, excessive increases are observed in June 2017 (Fig. 1, Fig. 2).

The series used in the model was compiled using the weekly data from January 2014/1 to December 2017/52 and the ARCH and GARCH methods. Firstly, the non-stationary series was made stationary by taking its logarithm and first difference (Fig. 1, Fig. 2). Secondly, it was found that inflation followed the AR(1) and MA(1) process by looking at the correlogram of ACF and PACF (Table 3). The modelling with AR(1) and MA(1) showed that both were significant. AIC and BIC values of the results obtained by running AR(1) and MA(1) individually were compared. As AIC and BIC values of AR(1) were lower than those of MA(1), AR(1) model was selected (Table 4, Table 5). Finally, the estimated ARCH/GARCH model was solved using the maximum likelihood (ML) technique, as shown in Table 4 and Table 5, and it was found that the model had the effect of GARCH (1, 1) and no effect of a higher degree of GARCH. The presence of the GARCH(1, 1) effect in the index series indicates that the GARCH variance series obtained as a result of the solution can be used as a measure of uncertainty in the veal cubes and minced meats series from January 2014/1 to December 2017/52.

The first difference of the data was taken to make the veal cubes series in Fig. 1 stationary. The series became stationary after taking its first difference.

The first difference of the data was taken to make the minced meats series in Fig. 2 stationary. The series became stationary after taking its first difference.

ADF and PP unit root tests indicate that the veal cubes series from January 2014/1 to December 2017/52 became stationary after taking its first difference. ADF and PP test values of the veal cubes series from January 2014/1 to December 2017/52 after taking its first difference are significant at a significance level of 1%, 5%
Determining the Change in Retail Prices of Veal

ADF and PP unit root tests indicate that the mince series from January 2014/1 to December 2017/52 became stationary after taking its first difference.

ADF and PP test values of the mince series from January 2014/1 to December 2017/52 after taking its first difference are significant at a significance level of 1%, 5% and 10%, respectively. The values in the parentheses are the optimal lag lengths.

In this framework, the model developed below for the mince series from January 2014/1 to December 2017/52 is significant.

\[
\sigma_t^2 = 0.087583 + 0.154908\sigma_{t-1}^2 - 0.647365\epsilon_{t-1}^2 \tag{3}
\]

The model indicates that the mince price is affected by the price in the previous week, with 1 unit of increase in the mince price being followed by a 0.15 unit of increase in the mince price in the next week.

\[
\sigma_t^2 = 0.140914 + 0.787804\sigma_{t-1}^2 - 0.133606\epsilon_{t-1}^2 \tag{4}
\]

The model indicates that the veal cubes price is affected by the price in the previous week, with 1 unit of increase in the veal cubes price being followed by a 0.78 unit of increase in the veal cubes price in the next week.

As an econometric result, it was found that the minced meats and veal cubes series from January 2014/1 to December 2017/52 had a quite varying variance.

In the models set on the prices of minced meats and veal cubes, RESID (-1) parameter shows the effect of shocks on price changes and GARCH (-1) parameter shows the effect of the previous period volatility on the current period volatility. The fact that the sum of both parameters is close to one means that it will have a long time effect on price volatility if a possible equilibrium situation deteriorates.

According to the results of the analysis, the veal cubes are more sensitive to the price changes compared to minced meats.

**DISCUSSION**

This study analysed the volatility of the retail prices of minced meats and veal cubes using volatility estimation modelling for the period between 2014 and 2017, and the results of the GARCH(1, 1) model show that it is the optimal model that best explains the change in the prices.

Analysis results indicate that the retail prices of minced meats and veal cubes saw excessive volatility in the period under consideration. Various policies and government interventions were implemented in the red meat sector in an effort to decrease the rising meat prices. In order to prevent price increases in October and November 2014, the General Directorate of Meat and Milk Board was authorised to import 15 thousand tons of beef with zero custom duty, and the custom duty for livestock imports...
by the private sector was reduced to 15%, provided that imported livestock would be used for stockbreeding [25]. Although the decisions taken were effective in reducing the retail prices of minced meats and veal cubes until the end of 2014, the prices started to elevate again as from early 2015 until May. The rising prices adversely affected the prices of foods and food services in this period [26]. In an attempt to prevent the price increases, first the custom duty for imports was reduced to 0%, and then the Meat and Milk Board was authorised to import 30,000 tons of fresh or frozen bovine meat [27,28].

Table 3. Correlogram results of the mince and dice series from January 2014/1 to December 2017/52

<table>
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<th>Type of Meat</th>
<th>Months</th>
<th>AC</th>
<th>PAC</th>
<th>Q Statistic of Error Terms</th>
<th>AC</th>
<th>PAC</th>
<th>Q Statistic of Error Terms</th>
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AC: Auto correlation, PAC: Partial auto correlation, Q-stat: Q statistic, Prob: Probability

To prevent the increase in red meat prices which were rising due to the lack of sufficient cattle for slaughter in Turkey, a market intervention was implemented through the Meat and Milk Board in March 2017, supplying butchers with fresh beef carcass and food and meat industrialists with frozen beef carcass through importation [29]. The upward movement in red meat prices observed since February 2017 continued in June, and the increase in prices was reported to be basically arising from the shortage of domestic supply [30]. Pursuant to the decision taken in response to the drastic price increase in June, the custom
Determining the Change in Retail Prices of Veal

The models developed for the time series data of the minced meats and veal cubes prices in the research show that increases in prices affect the next week's prices. The studies conducted on the same subject found that volatility of beef prices reacted heavily to the market movements, that a negative shock in the prices caused a higher volatility than a positive shock of the same degree, and that the fluctuations in beef prices resulted in fluctuations in the price of mutton and chicken.

Veal meat prices in Turkey were determined to be strictly affected from the economic and political instability, food crisis, natural events etc., and the prices of veal meat could return to normal values after a long time compared to the average price. On the other hand, the dependence on beef and veal red meat has increased over the years and it reflected to the importation as well. Red meat consumption has been sensitive to the price changes and always responded negatively to the beef meat pricing crisis. Red meat imports, unfortunately, couldn't help the price increases and the market price of red meat was increased later because of a lower local supply. On the other hand, due to the high level intermediary margins of beef meat marketing, red meat prices is increasing in Turkey.

In conclusion, the results of the estimation prove that the retail prices of veal cubes and minced meats fluctuated conspicuously in the period under consideration and that the implemented policies and the market interventions were not adequate to eliminate the instability and uncertainty of the prices. The presence of oligopsonistic market conditions, the long marketing chain in the sector and the import dependency to meet the demand for red meat attenuate the market power. Hence, it is essential that the price uncertainty is reduced from the perspective of producers and concrete steps other than importation are needed.
be taken to prevent increases in retail prices from the perspective of consumers so that the sector can grow and reach a level where it can compete globally.

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